

Lecture 3 - scribbles

Tuesday, September 10, 2019 14:29

- today:
 - finish prob. review
 - parametric models

Bayes rule

→ invert the conditioning

$$\text{by def. : } p(y|x) = \frac{p(x,y)}{p(x)}$$

Bayes rule

$$p(x|y) = \frac{p(y|x)p(x)}{p(y)} \sim \sum_x p(y|x)p(x)$$

$$\Rightarrow p(x,y) = p(y|x)p(x)$$

"product rule"

Chain rule : → successive application of product rule

$$\begin{aligned} p(x_1, \dots, x_n) &= p(x_n | x_{1:n-1}) p(x_{1:n-1}) \\ &\quad p(x_n | x_{1:n-1}) p(x_{n-1} | x_{1:n-2}) p(x_{1:n-2}) \\ &\quad \ddots \\ &= \prod_{i=1}^n p(x_i | x_{1:i-1}) \end{aligned}$$

Convention: $1:0 = \emptyset$, $p(x_i | x_\emptyset) = p(x_i)$

always true?
[can choose any
 $n!$ permutation]

→ can be simplified when making conditional independence assumptions
(from a directed graphical model)

Conditional independence

X is cond. indep. of Y given Z notation: $X \perp\!\!\!\perp Y | Z$

$$\Leftrightarrow p(x,y|z) = p(x|z)p(y|z) \quad \begin{array}{l} \forall x \in \mathcal{D}_X \\ \forall y \in \mathcal{D}_Y \\ \forall z \in \mathcal{D}_Z \text{ s.t. } p(z) \neq 0 \end{array}$$

Exercise to the reader: prove that if $X \perp\!\!\!\perp Y | Z$

$$\text{then } p(x|y,z) = p(x|z)$$

[conditional analog of $X \perp\!\!\!\perp Y \Rightarrow p(x|y) = p(x)$]

Example: Z = indicator whether mother carries a genetic disease

$X = \text{son 1 has disease}$

$X \perp\!\!\!\perp Y | Z$

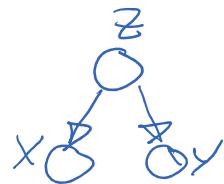
Example: Z - indicator whether mother carries a genetic disease

$$X = \text{son 1 has disease}$$

$$Y = \text{son 2 " "}$$

$$X \perp\!\!\!\perp Y | Z$$

here, X is not "marginally indep." of Y
(i.e. $X \perp\!\!\!\perp Y$)
 $\exists x, y \text{ s.t. } p(x, y) \neq p(x)p(y)$



$$p(x, y, z) = p(z|z)p(y|z)p(x|z)$$

Example of pairwise indep. R.V.'s that are not mutually independent

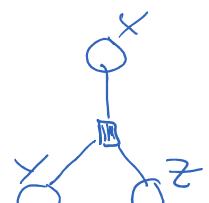
$$X = \text{coin flip as } \{0, 1\}$$

$$Y = \begin{matrix} \text{indep.} \\ \text{coin flip} \\ (X \perp\!\!\!\perp Y) \end{matrix}$$

$$\text{define } Z \triangleq X \text{ xor } Y$$

$$\begin{matrix} Z \perp\!\!\!\perp X \\ Z \perp\!\!\!\perp Y \end{matrix}$$

$$\cancel{Z \perp\!\!\!\perp (X, Y)}$$



(factor graph)

Other notation:

expectation/mean of a R.V.

$$\mathbb{E}[X] \triangleq \sum_{x \in \mathcal{X}} x p(x) \quad (\text{for discrete R.V.})$$

$$\int_{\Omega} x p(x) dx \quad (" \text{ cts. } ")$$

$$\mathbb{E}[\cdot] \text{ is a linear operator} \Rightarrow \mathbb{E}[aX + bY] = a\mathbb{E}[X] + b\mathbb{E}[Y]$$

scalar

Variance:

$$\text{Var}[X] \triangleq \mathbb{E}[(X - \mathbb{E}[X])^2]$$

"measure of dispersion"

$$= \mathbb{E}X^2 - (\mathbb{E}X)^2$$



$$\mathbb{E}[X] + \sigma = \sqrt{\text{Var}[X]}$$

"standard deviation"

parametric models:

↳ a family of distributions set of valid parameters

$$\mathcal{P}_{\Theta} = \left\{ p(x; \theta) \mid \theta \in \Theta \right\}$$

possible pmf/pdf
depending on parameter θ

abuse of notation: $\{ p(x; \theta) \mid \theta \in \Theta \}$

"better" notation $\{ p(X; \theta) \mid \theta \in \Theta \}$

from context:

"support of distribution" i.e. \mathcal{D}_X
is usually fixed for all θ

examples: $\mathcal{D}_X = \{0, 1\}$ for a coin flip

$\mathcal{D}_X = [0, \infty)$ for gamma
dist. family

Notation: $X \sim \text{Bern}(\theta)$

\hat{p} "R.V. X is distributed as a Bernoulli dist."

$$p(x; \theta) = \text{Bern}(x; \theta)$$

parameter with parameter θ
of dist.
variable for pmf

Bernoulli: prob. of a coin flip $\mathcal{D}_X = \{0, 1\}$

$$P\{X=1 \mid \theta\} = \theta \Rightarrow \Theta = [0, 1]$$

$$p(x; \theta) = \theta^x (1-\theta)^{1-x} = \text{Bern}(x; \theta)$$

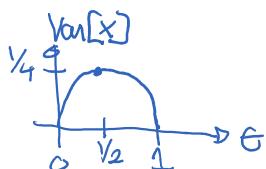
$$\mathbb{E}[X] = \theta$$

$$\text{Var}[X] = \theta(1-\theta)$$

$$\text{Var}[X; \theta]$$

$$\text{Var}_{\theta}[X]$$

$$\text{Var}[X \mid \theta]$$



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another example: $p(x; (\mu, \sigma^2)) = N(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$

$N(x; \mu, \sigma^2)$

binomial distribution: model n indep. coin flips

sum of n indep. $\text{Bern}(\theta)$ R.V.s

(note: θ is usually denoted by p)

let $X_i \sim \text{Bern}(\theta)$ ^{iid} \rightarrow independent & identically distributed

implicitly talking about X_1, \dots, X_n

Let $X = \sum_{i=1}^n X_i$; Then we have $X \sim \text{Bin}(n, \theta)$

"binomial with parameters $n \notin E$ "

$$\Omega_X = \{0, 1, \dots, n\}$$

$$p(x; \theta) = \binom{n}{x} \underbrace{\theta^x (1-\theta)^{n-x}}_{\substack{\# \text{ of ways} \\ \text{to choose } x \\ \text{elements out of } n \\ "n \text{ choose } x"}}, \quad \text{for } x \in \Omega_X = \{0, n\}$$

$$\binom{n}{x} \triangleq \frac{n!}{x!(n-x)!} \rightarrow \theta^{\sum x_i} (1-\theta)^{\sum (1-x_i)}$$

$$= \prod_{i=1}^n \text{Bern}(x_i; \theta) = p(x_1, \dots, x_n)$$

$$\text{mean: } X = \sum_i X_i$$

$$\mathbb{E}[X] = \sum_i \mathbb{E}[X_i] = n \theta$$

$$\text{Similarly } \text{Var}\left[\sum_i X_i\right] = \sum_i \text{Var}[X_i] = n \theta(1-\theta)$$

other distributions: Poisson(λ) $\Omega_X = \{0, 1, \dots\} = \mathbb{N}$ [count data]

Gaussian in 2D $\mathcal{N}(\mu, \sigma^2)$ mean variance

$\Omega_X = \mathbb{R}$

gamma $\Gamma(\alpha, \beta)$

shape scale

other: Laplace, Cauchy, exponential, beta, Dirichlet, etc.

statistical concepts

intuition

A

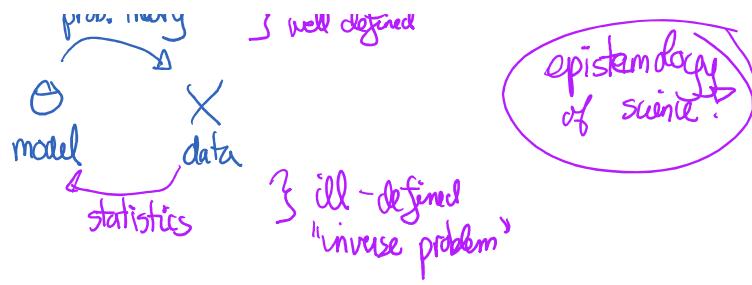
prob. theory

} well defined

epistemology
of science

Statistical concepts

Cartoon



example: model n indep. coin flips

prob. theory \rightarrow prob. K heads in a row

Statistics: I have observed K heads, what is θ ?
n-K tails

Frequentist vs. Bayesian:

Semantic of prob.: meaning of a prob.?

a) (traditional) frequentist semantic

$P\{X=x\}$ represents the limiting frequency of observing $X=x$

If I could repeat $\#$ of i.i.d. experiments

b) Bayesian (subjective) semantic

$P\{X=x\}$ encodes an agent "belief" that $X=x$

laws of prob. characterize a "rational" way to combine "beliefs"

and "evidence" [observations]

[\Rightarrow has motivation in terms of gambling, utility/decision theory, etc.]

operationally:

Bayesian approach: \textcircled{S} very simple philosophically:

- treat all uncertain quantities as R.V.

- i.e. encode all knowledge about the system ("beliefs")
as a prior on probabilistic model
and then use law of prob. (and Bayes rule) to get
answers!

frequentist justification:

• for a discrete R.V., suppose $P\{X=x\}=\theta$

$$\Rightarrow P\{X \neq x\} = 1-\theta$$

$$B \stackrel{\Delta}{=} \mathbf{1}\{X=x\} \Rightarrow B \sim \text{Bern}(\theta) \text{ R.V.}$$

indicator ft. $\mathbf{1}_A(u) = \begin{cases} 1 & \text{if } u \in A \\ 0 & \text{o.w.} \end{cases}$

repeat iid. experiments i.e. $B_i \stackrel{iid}{\sim} \text{Bern}(\theta)$

by L.L.N.
(law of large numbers) $\frac{1}{n} \sum_{i=1}^n B_i \xrightarrow{a.s.} \mathbb{E}[B_1] = \theta$

by CLT $\sqrt{n} \left(\frac{1}{n} \sum_{i=1}^n B_i - \theta \right) \xrightarrow{d} N(0, \theta(1-\theta))$